

EXTRA CHANCE IN MONTE CARLO

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New Perspectives in Markov Chain Monte Carlo
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INTRODUCTION & MOTIVATION



Markov chain Monte Carlo methods

Family of methods for the simulation of stochastic processes.

How to: given Π (up to \propto), Metropolis-Hasting generates a Markov chain $\{X_n\}_{n=0}^{\infty}$ for which Π is stationary.

Problem: correlation (the samples are not independent).

Solution: dynamics (gHMC) and, moreover, persistence (xHMC).



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Introduction & Motivation

gHMC & xHMC

Quantity & Quality

Case of Study & Benchmark

Epi & Logue



- Given Π (up to \propto), we define the potential energy

$$V(q) = -\log(\Pi(q))$$

- We consider the (fictional) Hamiltonian

$$H(q, p) = \frac{1}{2}p^T M^{-1}p + V(q)$$

- Instead of sampling $\Pi(q) = \exp(-V(q))$ in state-space $q \in \mathbb{R}^d$, we sample

$$\Pi(q, p) = \exp(-H(q, p)) = \exp\left(-\frac{1}{2}p^T M^{-1}p\right) \cdot \exp(-V(q))$$

in phase-space $(q, p) \in \mathbb{R}^d \times \mathbb{R}^d$.

Note: q and p are independently distributed, $p \sim N(0, M)$.

- Dynamics described by

$$H(q, p) = \frac{1}{2} p^T M^{-1} p + V(q)$$

- Motions characterized by

$$\dot{q} = M^{-1} p, \quad \dot{p} = -\nabla V$$

- Can be hard to integrate

$$(q(\tau), p(\tau)) = \mathbf{L}_\tau(q_0, p_0)$$

↪ numerical integrators

- Challenge: conservation laws

↪ energy, symplecticity (volume)

Verlet / Leapfrog integrator

$$p_{1/2} = p_0 + \frac{h}{2} f(q_0)$$

$$q_1 = q_0 + h p_{1/2}$$

$$p_1 = p_{1/2} + \frac{h}{2} f(q_1)$$

where $f = -\nabla V$

$$\mathbf{L}_\tau \rightsquigarrow \mathbf{L}_{s,h}$$

GHMC & xHMC



GENERALIZED HYBRID/HAMILTONIAN MONTE CARLO

```
01  $p_0 = 0$ 
02  $\xi_0 = (q_0, p_0)$ 
03 for  $n = 0$  to  $N - 1$  do
04    $\xi_n^0 = \mathbf{R}\xi_n$ 
05    $\xi_n^1 = \mathbf{L}\xi_n^0$ 
06    $\Sigma = \min(1, \Pi(\xi_n^1)/\Pi(\xi_n^0))$ 
07   if  $u_n \leq \Sigma$  then
08      $\xi_{n+1} = \xi_n^1$ 
09   else
10      $\xi_{n+1} = \mathbf{F}\xi_n^0$ 
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Step 0: preparation

- Space extension

$$q \implies \xi = (q, p)$$

- Hamiltonian dynamics

$$H(\xi) = \frac{1}{2}p^T M^{-1}p + V(q)$$

- Maxwell-Boltzmann distr.

$$\Pi(\xi) \propto \exp(-H(\xi))$$

- Start the loop

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Step 1: momentum refreshment

$$\begin{aligned} \mathbf{R}\xi &= \mathbf{R}_\psi(q, p) := (q, p') \\ p' &:= \cos \psi \cdot p + \sin \psi \cdot z \\ z &\sim N(0, 1) \end{aligned}$$

where $\psi \in [0, \pi/2] \rightsquigarrow \sin \psi \in [0, 1]$

Note: if $\sin \psi = 1$ then

$$p' = p(\mathbf{R}\xi) \sim N(0, 1)$$

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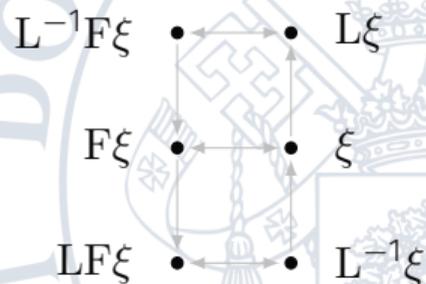
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Step 2: integrate & flip

$$\mathbf{L}: (q_0, p_0) \xrightarrow{L_{s,h}} (q_1, p_1)$$

$$\mathbf{F}: (q, p) \mapsto (q, -p)$$



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Step 3: accept/reject

$$\xi_{n+1} = \begin{cases} \mathbf{L}\xi_n & \text{w. prob } \Sigma \\ \mathbf{F}\xi_n & \text{w. prob } 1 - \Sigma \end{cases}$$

by drawing $u_n \sim U([0, 1])$

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Step 4: throw momentum away



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Finale: genHMC algorithm



Goal: improve acceptance rate (and chain quality).

Problem: rejections are costly.

Question: can we salvage them?



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Idea: generate an integration chain and choose (according to the dynamics)

$K \geq 1$ total chances
 $K - 1 \geq 0$ extra chances



 J. Sohl-Dickstein, M. Mudigonda & M.R. DeWeese
Hamiltonian Monte Carlo Without Detailed Balance.
Proc. Int. Conf. Mach. Learn. **31** (2014).

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Problem: rejections are costly.

Question: can we salvage them?

$$\xi_{n+1} = \begin{cases} L^1 \xi_n & \text{with probability } \pi_{L^1}(\xi_n) \\ L^2 \xi_n & \text{with probability } \pi_{L^2}(\xi_n) \\ \vdots & \\ L^K \xi_n & \text{with probability } \pi_{L^K}(\xi_n) \\ F \xi_n & \text{with probability } \pi_F(\xi_n) \end{cases}$$

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$$\begin{aligned}\Sigma^a(\xi) &:= \min \left\{ 1, \max_{b < a} \frac{\Pi(\text{FL}^b \xi)}{\Pi(\xi)} \right\} \\ \pi_{L^{a+1}}(\xi) &:= \Sigma^{a+1}(\xi) - \Sigma^a(\xi) \\ \pi_F(\xi) &:= 1 - \Sigma^K\end{aligned}$$

Theorem (Detailed Balance [Campos & Sanz-Serna, 2015])

The associated transition kernel satisfies detailed balance.

Proof.

Thanks to reversibility and volume preservation, establish $\Pi(\xi) \cdot \pi_{L^a}(\xi) = \Pi(\text{FL}^a \xi) \cdot \pi_{L^a}(\text{FL}^a \xi)$. Then, go for the definition detailed balance. \square

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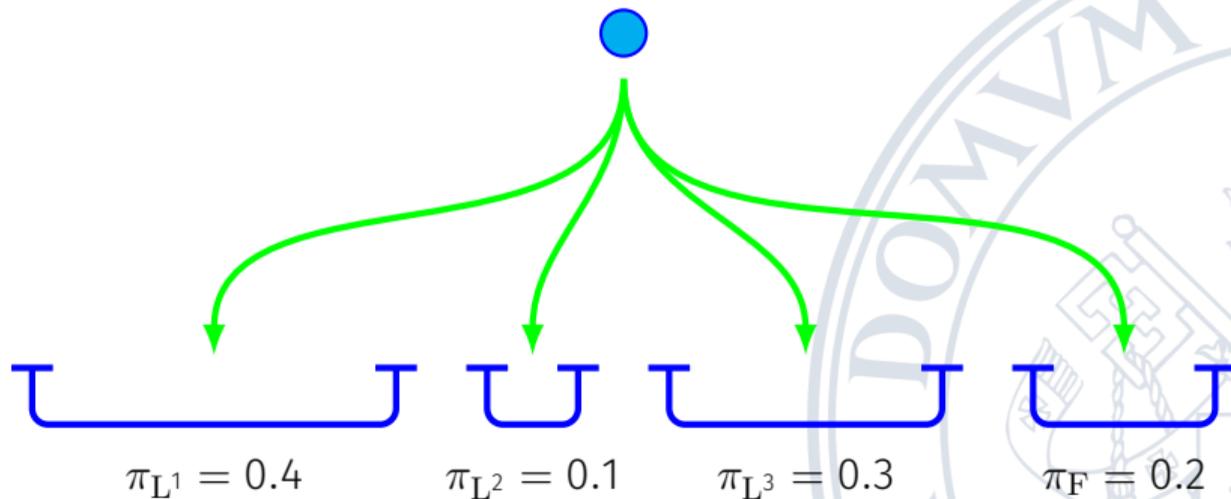
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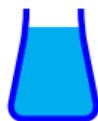
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$u = 0.6$



$\pi_{L^1} = 0.4$



$\pi_{L^2} = 0.1$

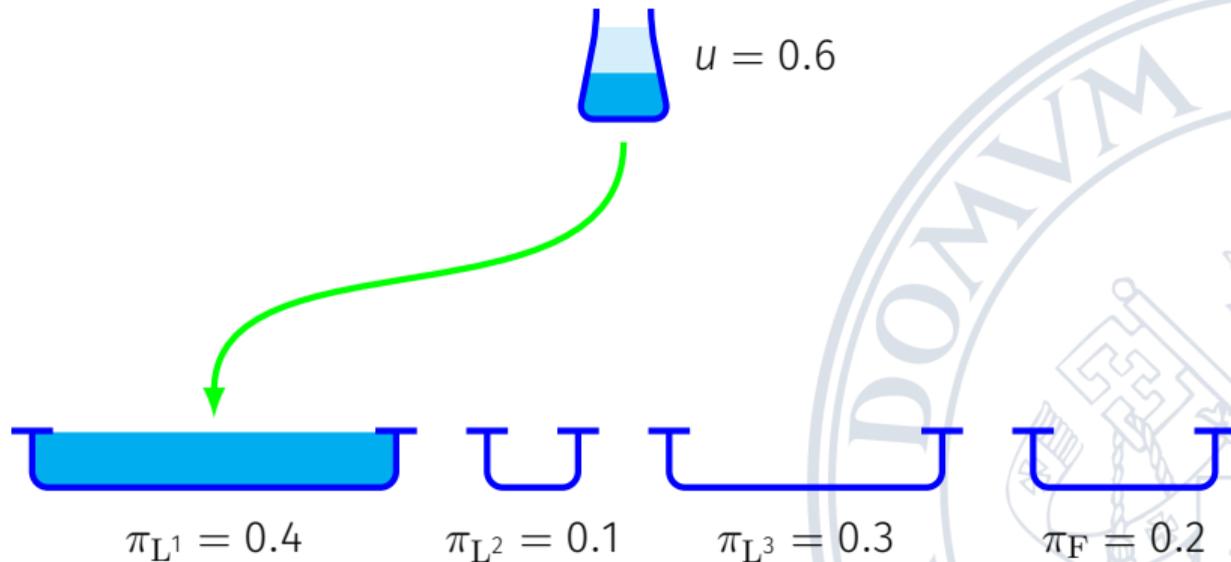


$\pi_{L^3} = 0.3$

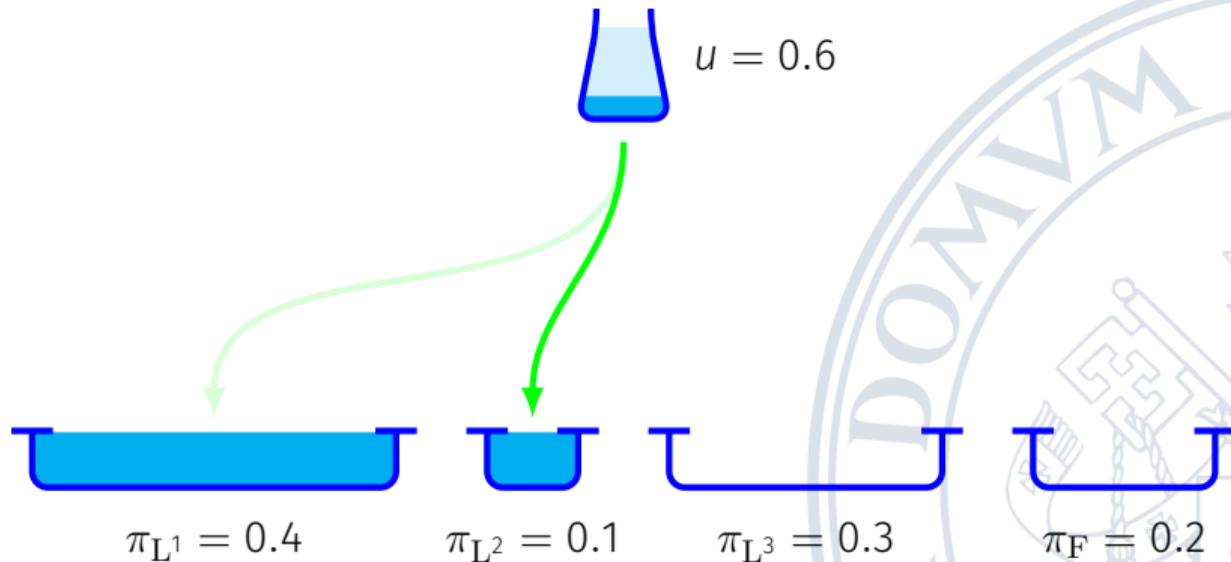


$\pi_F = 0.2$

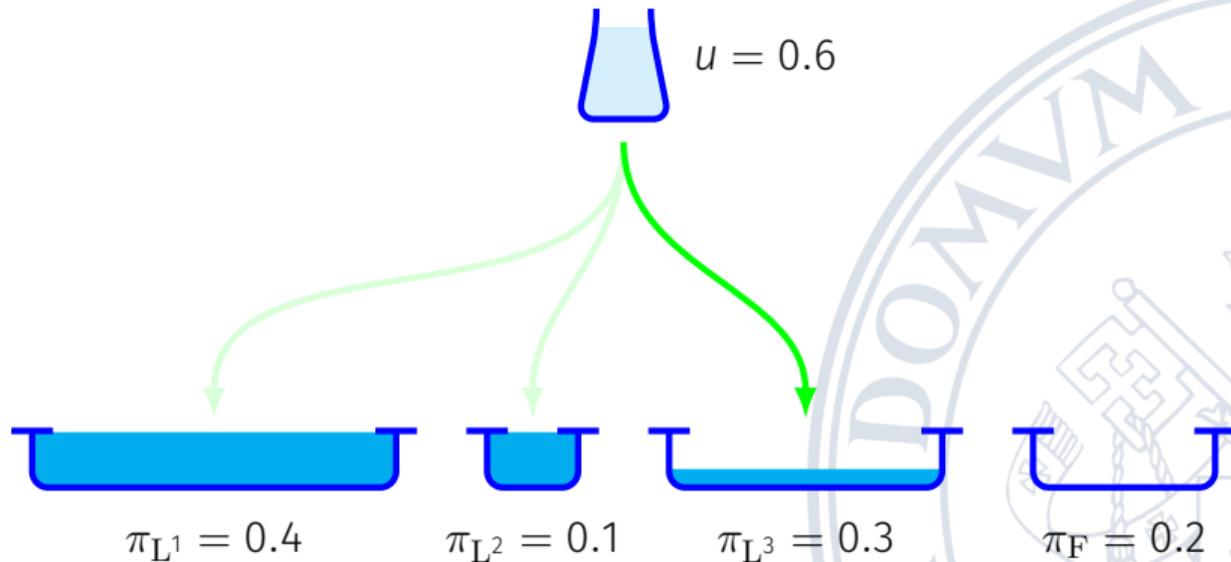
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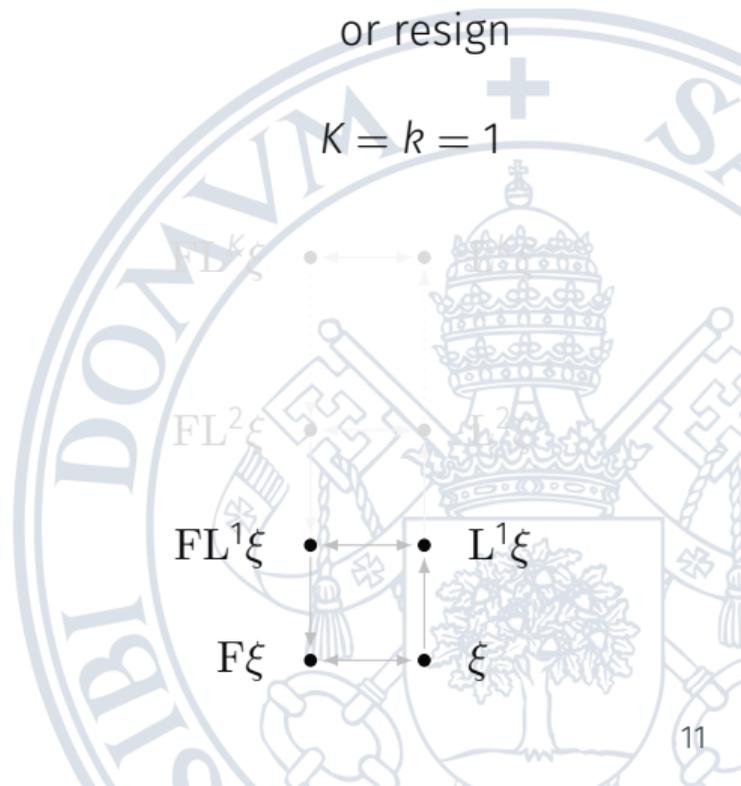


THE xHMC ALGORITHM

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05    $\Sigma = k = 0$ 
06   do
07      $k = k + 1$ 
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genHMC: succeed
or resign

$K = k = 1$



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xHMC: persist

$K \geq k \geq 1$

$\text{FL}^k \xi \longleftrightarrow \text{L}^k \xi$

$\text{FL}^2 \xi \longleftrightarrow \text{L}^2 \xi$

$\text{FL}^1 \xi \longleftrightarrow \text{L}^1 \xi$

$\text{F}\xi \longleftrightarrow \xi$

QUANTITY & QUALITY



EFFECTIVE SAMPLE SIZE (ESS)

How to quantitatively determine the meaningful data of an MC?

X_1, X_2, \dots, X_m MC $\sim \Pi$

$$\mu_m^X \rightarrow \mu$$

$$\sqrt{m}(\mu_m^X - \mu) \rightarrow N(0, \sigma^2)$$

$$\mu = \int g(x)\Pi(x) dx$$

$$\mu_r^Z = \frac{1}{r} \sum_{k=1}^r g(Z_k)$$

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$$\frac{\sigma^2}{m} \approx \frac{s^2}{n}$$

\Rightarrow

$$n \approx m \frac{s^2}{\sigma^2}$$

$$\sigma^2 \approx ?_i$$

$$s^2 \approx \frac{1}{m} \sum_{i=1}^m (g(X_i) - \mu_m^X)^2 =: \gamma_{m,0}^X$$

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$$\mu_m^X \rightarrow \mu$$
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$$\mu = \int g(x)\Pi(x) dx$$
$$\mu_r^Z = \frac{1}{r} \sum_{k=1}^r g(Z_k)$$

$$Y_1, Y_2, \dots, Y_n \text{ i.i.d.} \sim \Pi$$
$$\mu_n^Y \rightarrow \mu$$
$$\sqrt{n}(\mu_n^Y - \mu) \rightarrow N(0, s^2)$$

Theorem (Kipnis & Varadhan, 1986)

For a stationary, irreducible, reversible Markov chain

$$m \text{ var } \mu_m^X \rightarrow \sigma^2 := \sum_{-\infty}^{+\infty} \gamma_k, \quad \gamma_k := \text{cov}(g(X_i), g(X_{i+k})) .$$

If σ^2 is finite, then

$$\sqrt{m}(\mu_m^X - \mu) \rightarrow N(0, \sigma^2)$$

EFFECTIVE SAMPLE SIZE (ESS)

- Direct approaches to compute γ_k give bad results.
- Alternative, define $\Gamma_k := \gamma_{2k} + \gamma_{2k+1}$

Theorem (Geyer, 1992)

$\Gamma_k = \Gamma(k)$ is strictly positive, strictly decreasing and strictly convex.

- Approximate $\gamma_k \approx \gamma_{m,k}^X := \frac{1}{m} \sum_{i=1}^{m-k} (g(X_i) - \mu_m^X)(g(X_{i+k}) - \mu_m^X)$
- and $\Gamma_k \approx \Gamma_{m,k}^X := \gamma_{m,2k}^X + \gamma_{m,2k+1}^X$
- Modify $\Gamma_{m,k}^X$ such that theorem is “satisfied”.
- Then $\sigma^2 \approx -\gamma_{m,0}^X + 2 \sum \Gamma_{m,k}^X$.



C.J. Geyer

Practical Markov Chain Monte Carlo.
Statistical Science 7 (1992).

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CASE OF STUDY & BENCHMARK



Linear alkanes are chemical compounds of the form



The conformation of the molecule is characterized by the positions of the Carbon atoms.



E. Cancès, F. Legoll & G. Stoltz

Theoretical and numerical comparison of some sampling methods for molecular dynamics.

ESAIM: M2AN **41** (2007).

It is assumed they behave according to the potential

$$V(q) = \sum_{i=1}^{n-1} V_2(d_{i,i+1}) + \sum_{i=1}^{n-2} V_3(\theta_i) + \sum_{i=1}^{n-3} V_4(\phi_i) + \sum_{i=1}^{n-4} \sum_{j=i+3}^n V_{LJ}(d_{i,j})$$

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where

$$V_2(d) = \frac{1}{2}k_0(d - d_0)^2$$

covalent bond potential between two consecutive Carbon atoms

- d , distance between the atoms
- k_0, d_0 , system parameters

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where

$$V_3(\theta) = \frac{1}{2} \kappa_0 (\theta - \theta_0)^2$$

three-body interaction potential

- θ , bending angle between three consecutive atoms
- κ_0, θ_0 , system parameters

Linear alkanes are chemical compounds of the form



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where

$$V_4(\phi) = c_1(1 - x) + 2c_2(1 - x^2) + c_3(1 + 3x - 4x^3)$$

four-body interaction potential

- $\phi = \arccos(x)$, dihedral angle between four consecutive atoms
- c_1, c_2, c_3 , system parameters

Linear alkanes are chemical compounds of the form



It is assumed they behave according to the potential

$$V(q) = \sum_{i=1}^{n-1} V_2(d_{i,i+1}) + \sum_{i=1}^{n-2} V_3(\theta_i) + \sum_{i=1}^{n-3} V_4(\phi_i) + \sum_{i=1}^{n-4} \sum_{j=i+3}^n V_{LJ}(d_{i,j})$$

where

$$V_{LJ}(d) = 4\epsilon \left(\left(\frac{\sigma}{d} \right)^{12} - \left(\frac{\sigma}{d} \right)^6 \right)$$

Lennard-Jones potential of two Carbon atoms connected by three or more covalent bonds

- d , distance between the atoms
- ϵ, σ , system parameters

SIMULATIONS

- alkane = 9 carbons (i.e. 6 diedra)
- $K = 0$, 3 (extra chances)
- $\sin \psi = 0.10, 0.25, 0.50, 0.75, 1$
- $L\Delta t = 0.12, 0.24, 0.36, 0.48, 0.60, 0.72$
- $\Delta t = 0.012, 0.020, 0.024, 0.030$
- chains per setting = 10

total chains = 1800 (several GiB)

- cost = 10^6 force evaluations
- ESS measures:
 - 1st diedra, $g(x) = \mathbf{1}_{[-1.75, 1.75]}(x)$

92.6707	9240	3368	3632
77.0985	8853	3782	3965
65.3370	7431	3401	3518
92.9657	5.8980	0.9573	0.1426
80.2533	15.9591	3.2026	0.4637
71.6308	21.9310	5.3675	0.8597
92.5993	8943	3436	3504
77.1728	8822	4160	4096
65.2272	8421	4352	4146
92.8594	6.0066	0.9512	0.1486
80.0717	16.0101	3.3024	0.4854
71.4885	22.0166	5.4070	0.8759
92.5784	9952	4689	4510
77.2683	10530	5278	5117
65.3837	9445	4997	4745
92.9053	5.9769	0.9382	0.1476
79.9234	16.1746	3.2781	0.4934
71.2970	22.0998	5.5346	0.8526
92.5878	8169	4135	3961
77.2775	9736	5410	5231
65.5464	8982	5168	4937
92.8834	5.9911	0.9473	0.1382
79.9234	16.1746	3.2781	0.4934
71.2970	22.0998	5.5346	0.8526
92.5878	8169	4135	3961
77.2775	9736	5410	5231
65.5464	8982	5168	4937
92.8834	5.9911	0.9473	0.1382

RESULTS: ACCEPTANCE RATIO

Δt	$K = 0$	$K = 3$				
		a_0	a_1	a_2	a_3	$\sum a_k$
0.012	93%	93%	6%	1%	0%	100%
0.016	86%	87%	11%	2%	0%	100%
0.020	77%	80%	16%	3%	1%	100%
0.024	65%	71%	22%	6%	1%	100%

Acceptance rates (rounded) for $\sin \psi = 1$, $L\Delta t = 0.48$ and different values of Δt and K , $K = 0$ (standard HMC) and $K = 3$ (three extra chances). The column labelled a_k gives the acceptance rate at the $(k + 1)$ -try.

Conclusion: extra chance increases acceptance ratio

Question: worth? (tricky, shorter chain lengths)

RESULTS: ACCEPTANCE RATIO

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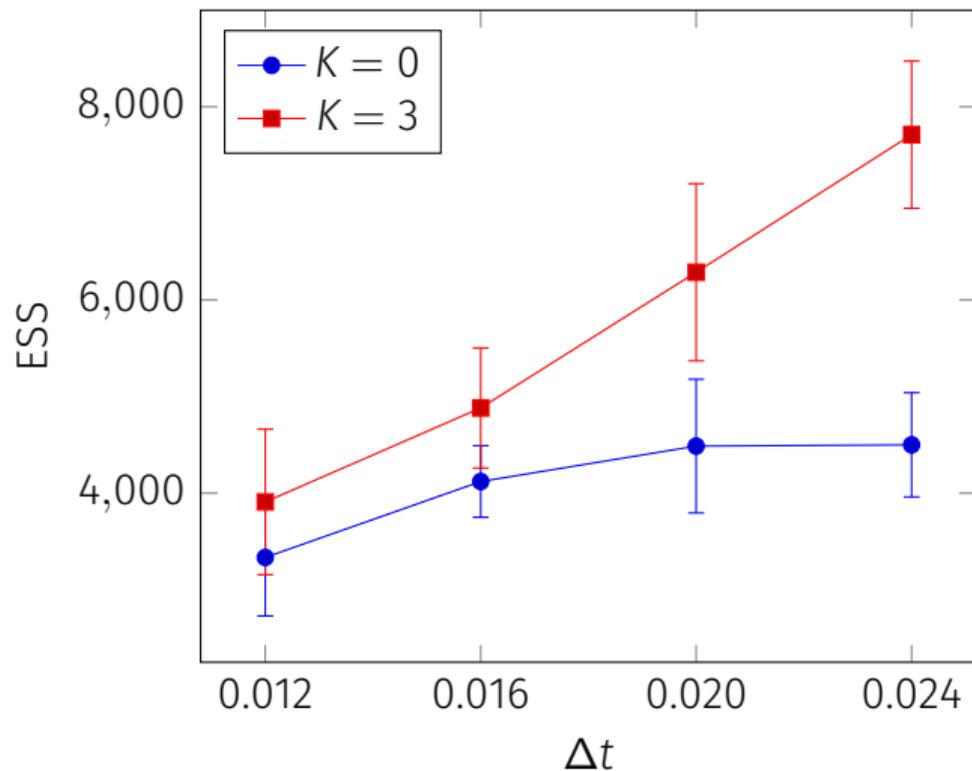
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RESULTS: EFFECTIVE SAMPLE SIZE VS. INTEGRATION TIME STEP

$$L\Delta t = 0.48, \sin \psi = 1$$

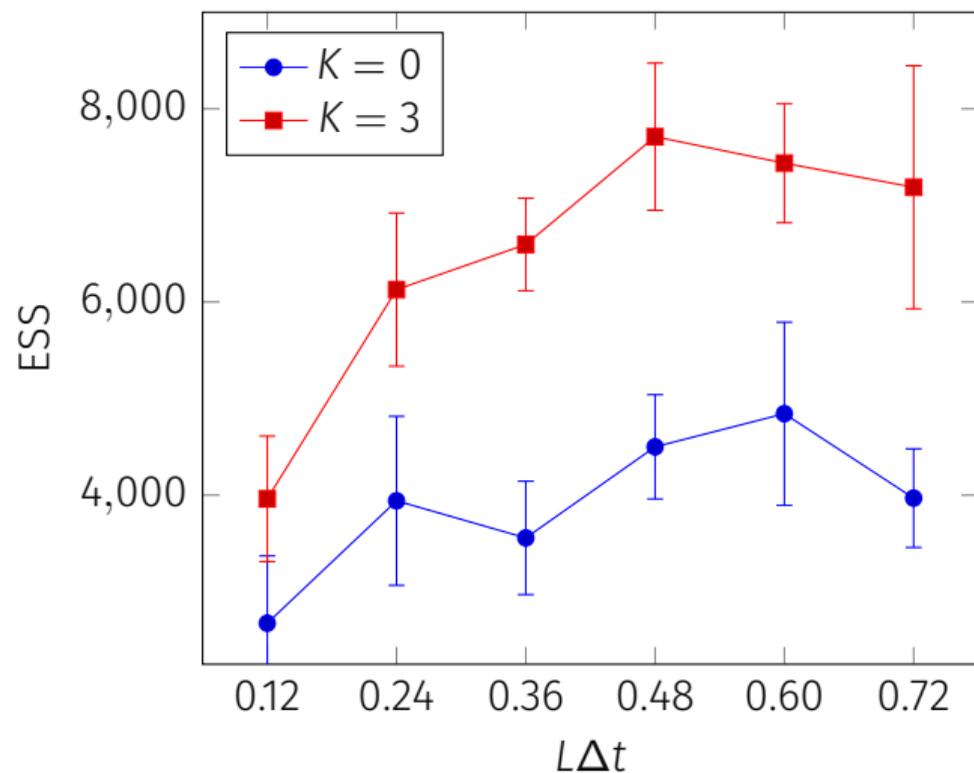


ESS as a function of ITS

- fixed ITI
- full mom. refresh

RESULTS: EFFECTIVE SAMPLE SIZE VS. INTEGRATION TIME INTEGRATION

$\Delta t = 0.024, \sin \psi = 1$

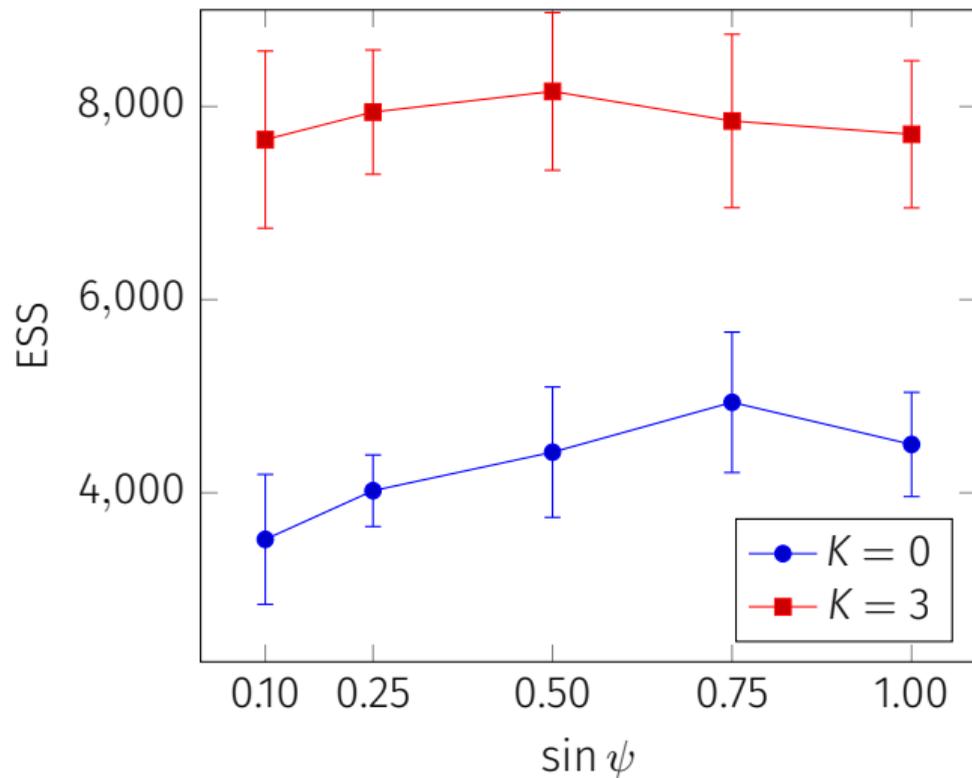


ESS as a function of ITI

- fixed ITS
- full momentum refreshment

RESULTS: EFFECTIVE SAMPLE SIZE VS. MOMENTUM REFRESHMENT

$\Delta t = 0.024, L\Delta t = 0.48$



ESS as a function of refreshment

- fixed ITS
- fixed ITI

EPI & LOGUE



Conclusions:

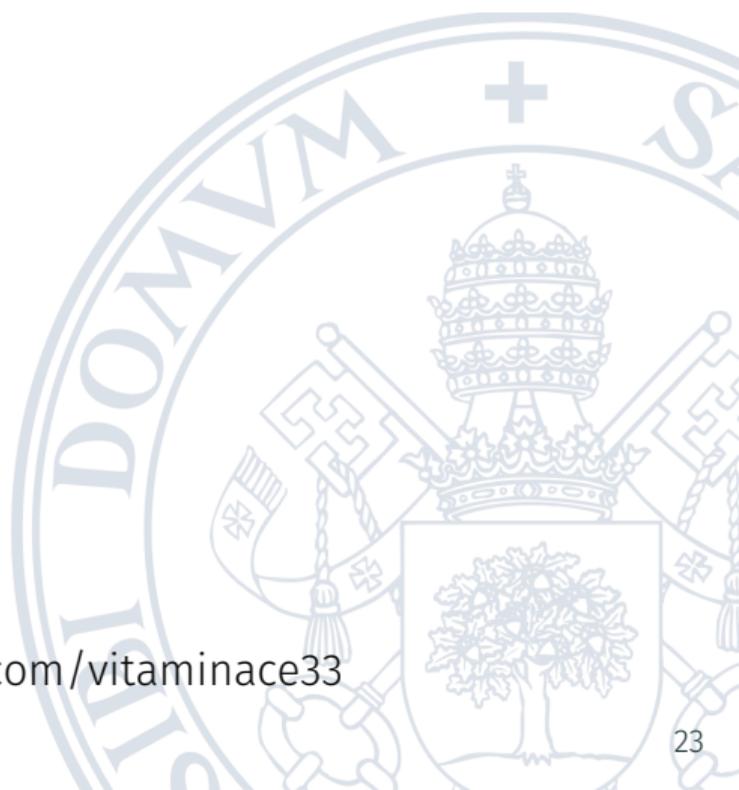
- clever (uses dynamics)
- applicable (detailed balance)
- easily implementable (simple back-recurrence)
- worth (quality chains)

Outlook:

- integration (cheap & tailored)

Extra:

- Matlab/Octave code available at <http://github.com/vitaminace33>



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Proc. Int. Conf. Mach. Learn. **31** (2014).

QUESTIONS?

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